

Two-stage Bayesian variable selection for linear models using I-priors

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Outline

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- Bayesian linear regression

- Types of priors

② ASIDE: Regression modelling using I-priors

- Linear regression and motivation for I-priors

- Examples of I-prior models

- I-prior summary

③ Bayesian variable selection

- Introduction

- Bayesian variable selection methods

④ Using I-priors in Bayesian variable selection

- Variable selection with I-priors

- Simulation study

- Two-stage procedure

- g-priors

⑤ Summary

Bayesian linear regression

- Consider a linear regression model for n observations on p variables:

$$\begin{aligned}\mathbf{y} &= \alpha + \mathbf{X}\beta + \epsilon \\ \epsilon &\sim \mathcal{N}(\mathbf{0}, \psi^{-1}\mathbf{I}_n)\end{aligned}\tag{1}$$

where $\alpha = \alpha\mathbf{1}_n$.

- The OLS estimate for β is $\hat{\beta} = (\mathbf{X}^T\mathbf{X})^{-1}\mathbf{X}^T\mathbf{y}$.
- The Bayesian approach supplements the data with additional information in the form of prior beliefs about the parameters:
 - ▶ $\alpha \sim \mathcal{N}(a, b)$
 - ▶ $\beta \sim \mathcal{N}(\mathbf{c}, \mathbf{D})$
 - ▶ $\psi \sim \Gamma(e, f)$
- Inference on the parameters Θ is through the posterior

$$f(\Theta|\mathbf{y}) \propto \overbrace{f(\mathbf{y}|\Theta)}^{\text{likelihood}} \times \overbrace{f(\Theta)}^{\text{prior}}$$

Types of priors and the I-prior

- Priors can either be pure beliefs (subjective) or chosen according to some principle (objective). Either way, they can also be
 - ▶ Informative - has an impact on the results
 - ▶ Uninformative - provides little or vague information
 - ▶ Improper - may not be a proper distribution

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- **I-priors (for regression coefficients)**

An I-prior on β for the linear model in (1) is a distribution on β such that its covariance matrix is the Fisher information of β . Also, assign a “best guess” on the prior mean, e.g. $\beta_0 = \mathbf{0}$.

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- An objective and information theoretic prior for linear models with an intuitive appeal:

\uparrow Fisher information $\Rightarrow \uparrow$ variance $\Rightarrow \downarrow$ influence of prior mean.

The I-prior linear regression model

$$\mathbf{y} = \alpha + \mathbf{X}\beta + \epsilon$$

$$\epsilon \sim N(\mathbf{0}, \psi^{-1} \mathbf{I}_n)$$

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$$\beta \sim N(\mathbf{0}, \lambda^2 \psi \mathbf{X}^T \mathbf{X}).$$

- λ is introduced to resolve the scale of measurements of \mathbf{X} .
- Assumption: *All variables are measured on the same scale, or at least standardised. More on this later...*

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- λ is introduced to resolve the scale of measurements of \mathbf{X} .
- Assumption: *All variables are measured on the same scale, or at least standardised. More on this later...*
- To complete the Bayesian model specification, set priors on the intercept and precision

$$\alpha \sim N(0, 1000)$$

$$\psi \sim \Gamma(0.001, 0.001).$$

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⑤ Summary

Linear regression

- Definition (**The linear regression model**)

$$y_i = f(x_i) + \epsilon_i$$

$y_i \in \mathbb{R}$, real-valued observations

$x_i \in \mathcal{X}$, a set of characteristics for unit i (2)

$f \in \mathcal{F}$, a vector space of functions over the set \mathcal{X}

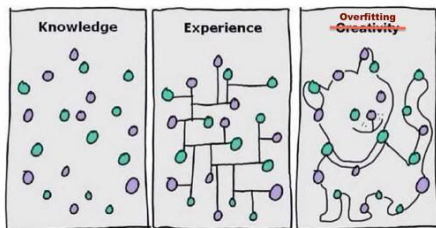
$$(\epsilon_1, \dots, \epsilon_n) \sim N(\mathbf{0}, \Psi^{-1})$$

$$i = 1, \dots, n$$

Note: For iid observations, $\Psi = \psi \mathbf{I}_n$. In general, $\Psi = (\psi_{ij})$.

Motivation for I-priors: The issue of overfitting

- When dimensionality is large, maximum likelihood overfits. Solutions:
 - ▶ Dimension reduction
 - ▶ Random effects models
 - ▶ Regularization...all require additional assumptions.
- I-priors require no assumptions other than those pertaining to the model of interest.
- But we do need a structural requirement for \mathcal{F} in the form of an inner-product space (reproducing kernel Hilbert/Krein space).



credits: <http://blog.sciencenet.cn/u/jerrycueb>

Functional vector spaces

Inner products

Kernel methods

Reproducing kernels

Hilbert spaces

Gaussian random vectors

I-prior theory

Fisher Information

Krein spaces

Means of random functions

Feature maps

Variances of random functions

Moore-Aronszajn Theorem

Random functions

Definitions & theorem

- Theorem (**Gaussian I-priors**) [Bergsma, 2014]

For the linear regression model (2), let \mathcal{F} be the RKKS with kernel $h : \mathcal{X} \times \mathcal{X} \rightarrow \mathbb{R}$. Then, assuming it exists, the Fisher information for f is given by

$$I[f](x_i, x'_i) = \sum_{k=1}^n \sum_{l=1}^n \psi_{kl} h(x_i, x_k) h(x'_i, x_l).$$

Let π be a Gaussian I-prior on f with prior mean f_0 and variance $I[f]$. Then π is called an I-prior for f , and a random vector $f \sim \pi$ has the random effect representation

$$f(x_i) = f_0(x_i) + \sum_{k=1}^n h(x_i, x_k) w_k$$

$$(w_1, \dots, w_n) \sim N(\mathbf{0}, \Psi).$$

Back to the (standard) linear regression model

$$\mathbf{y} = \alpha + \mathbf{X}\beta + \epsilon$$

$$\epsilon \sim N(\mathbf{0}, \psi^{-1} \mathbf{I}_n)$$

$$\beta \sim N(\beta_0, \lambda^2 \psi \mathbf{X}^T \mathbf{X})$$

Back to the (standard) linear regression model

$$\begin{aligned}\mathbf{y} &= \overbrace{\boldsymbol{\alpha} + \mathbf{X}\boldsymbol{\beta}}^{\mathbf{f}} + \boldsymbol{\epsilon} \\ \boldsymbol{\epsilon} &\sim N(\mathbf{0}, \psi^{-1}\mathbf{I}_n) \\ \boldsymbol{\beta} &\sim N(\boldsymbol{\beta}_0, \lambda^2\psi\mathbf{X}^T\mathbf{X})\end{aligned}$$

Equivalently,

$$\begin{aligned}\boldsymbol{\beta} &= \boldsymbol{\beta}_0 + \lambda\mathbf{X}^T\mathbf{w} \\ \mathbf{w} &\sim N(\mathbf{0}, \psi\mathbf{I}_n).\end{aligned}$$

Thus, an I-prior on \mathbf{f} is

$$\begin{aligned}\mathbf{f} &= \overbrace{\boldsymbol{\alpha} + \mathbf{X}\boldsymbol{\beta}_0}^{\mathbf{f}_0} + \overbrace{\lambda\mathbf{X}\mathbf{X}^T\mathbf{w}}^{\mathbf{H}_\lambda\mathbf{w}} \\ \mathbf{w} &\sim N(\mathbf{0}, \psi\mathbf{I}_n).\end{aligned}$$

Toolbox of RKHS/RKKS

- Choose different $\{\mathcal{F}, h\}$ to suit type of data to model.

$\mathcal{X} = \{x_i\}$	Characteristic/Uses	Vector space \mathcal{F}	Kernel $h(x_i, x_k)$
Nominal	1) Categorical covariates; 2) In a multilevel setting, x_i = group no. of unit i .	Pearson	$\frac{\mathbb{I}[x_i=x_k]}{p_i} - 1$ where $p_i = \mathbb{P}[X = x_i]$
Real	As in classical regression, x_i = real-valued covariate associated with unit i .	Canonical	$x_i x_k$
Real	As in (1-dim) smoothing, x_i = data point associated with observation y_i .	Fractional Brownian Motion (FBM)	$ x_i ^{2\gamma} + x_k ^{2\gamma} - x_i - x_k ^{2\gamma}$ with $\gamma \in (0, 1)$
Nominal + Real	Used for random intercept/slope modelling.	ANOVA	Pearson + Canonical kernels

Example: Simple linear regression

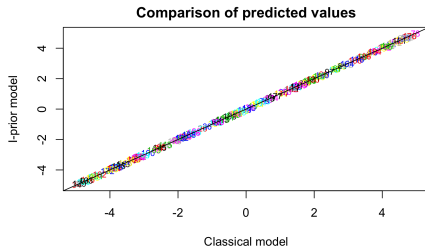
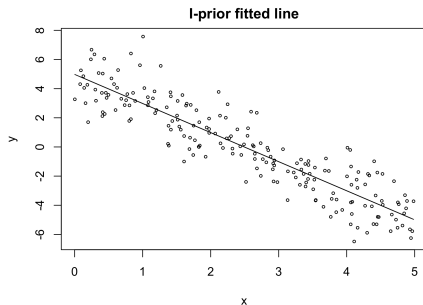
Classical model

$$y_i = \beta_0 + \beta_1 x_i + \epsilon_i$$
$$\epsilon_i \sim N(0, \sigma^2)$$

I-prior model

$$y_i = \alpha + \sum_{k=1}^n h_{\lambda}(x_i, x_k) w_k + \epsilon_i$$
$$\epsilon_i \sim N(0, \psi^{-1})$$
$$w_i \sim N(0, \psi)$$

h_{λ} is the Canonical kernel



$$\text{MSE}(\text{classical}) = 1.770 \quad \text{MSE}(\text{I-prior}) = 1.770$$

Example: 1-dimensional smoothing

Classical model

$$y_i = \beta_0 + \beta_1 x_i + \beta_2 x_i^2 + \beta_3 x_i^3$$

$$\epsilon_i \sim N(0, \sigma^2)$$

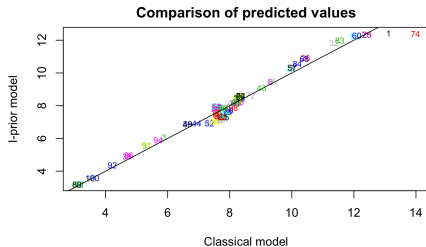
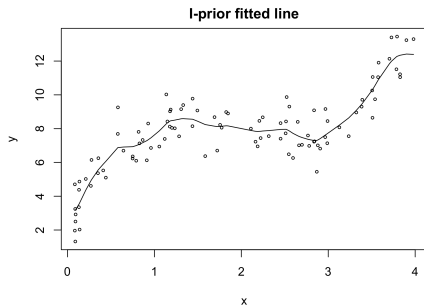
I-prior model

$$y_i = \alpha + \sum_{k=1}^n h_{\lambda, \gamma}(x_i, x_k) w_k + \epsilon_i$$

$$\epsilon_i \sim N(0, \psi^{-1})$$

$$w_i \sim N(0, \psi)$$

$h_{\lambda, \gamma}$ is the FBM kernel



$$\text{MSE}(\text{classical}) = 0.987 \quad \text{MSE}(\text{I-prior}) = 0.836$$

Example: Multilevel modelling

Classical model

$$y_{ij} = \beta_{0j} + \beta_{1j}x_{ij} + \epsilon_{ij}$$

$$\begin{pmatrix} \beta_{0j} \\ \beta_{1j} \end{pmatrix} \sim N\left(\begin{pmatrix} \beta_0 \\ \beta_1 \end{pmatrix}, \begin{pmatrix} \phi_0 & \phi_{01} \\ \phi_{01} & \phi_1 \end{pmatrix}\right)$$

$$\epsilon_{ij} \sim N(0, \sigma^2)$$

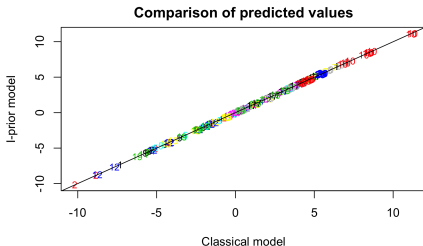
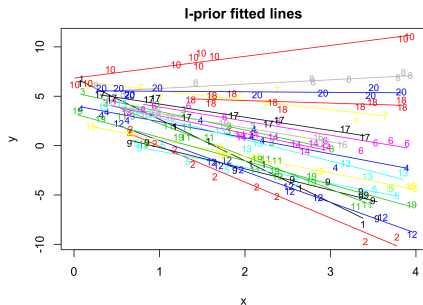
I-prior model

$$y_i = \alpha + \sum_{k=1}^n h_{\lambda}(x_i, x_k) w_k + \epsilon_i$$

$$\epsilon_i \sim N(0, \psi^{-1})$$

$$w_i \sim N(0, \psi)$$

h_{λ} is the ANOVA kernel



$$\text{MSE(classical)} = 0.227 \quad \text{MSE(I-prior)} = 0.226$$

I-prior summary

- The I-prior methodology is a modelling technique that guards against overfitting linear models when dimensionality is large relative to sample size, with advantages such as
 - ▶ Model parsimony
 - ▶ Requires no additional assumptions
 - ▶ Simpler estimation (EM algorithm)
- Many models shown to work with using I-priors such as multiple regression, smoothing models, random effects models and growth curve models.
- Areas of research include
 - ▶ Extension to GLMs
 - ▶ Structural Equation Models
 - ▶ Models with structured error covariances
- Key idea: *Fisher information as the covariance matrix for priors.*

① Introduction

② ASIDE: Regression modelling using I-priors

③ Bayesian variable selection

④ Using I-priors in Bayesian variable selection

⑤ Summary

Model selection criteria

- Would like to search the entire model space to find the “best” model based on a certain criterion.
- Many methods for model selection criteria... (adjusted) R^2 , AIC, BIC, Mallows's C_p , (k -fold) cross-validation, posterior model odds, Bayes factors, etc.
- When a large set of models to be compared, most tasks can be computationally prohibitive or even unfeasible.

Bayesian model evaluation

- It is believed that a set of data \mathbf{Y} has been generated from the pdf $f(\mathbf{y}|m_k, \boldsymbol{\Theta}_k)$, where m_k is one of a set of $M = \{m_1, \dots, m_K\}$ models.

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- As Bayesians do...
 - ▶ Assign priors $f(\Theta_k|m_k)$ and $f(m_k)$
 - ▶ Compute the posterior

$$\begin{aligned}
 f(m_k|\mathbf{y}) &\propto f(\mathbf{y}|m_k)f(m_k) \\
 &\propto \int f(\mathbf{y}|m_k, \Theta_k)f(\Theta_k|m_k) d\Theta_k f(m_k)
 \end{aligned}$$

- ▶ Choose m_k with highest posterior probability

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 \end{aligned}$$

- Choose m_k with highest posterior probability
- Use MCMC methods to sample from the posterior when
 - the integral in the posterior is not analytically tractable; and/or
 - the model space is too large to make any calculation of the posterior for all models unfeasible.

Bayesian variable selection

- Consider again the linear model in (1)

$$y_i = \alpha + \beta_1 X_{i,1} + \cdots + \beta_p X_{i,p} + \epsilon_i$$

$$\epsilon_i \sim N(0, \psi^{-1}) \text{ iid}$$

$$i = 1, \dots, n$$

- A model is a subset of variables $\{\tilde{X}_1, \dots, \tilde{X}_q\}$ from $\{X_1, \dots, X_p\}$. There are 2^p models to consider.

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- Index each of these 2^p models by the vector

$$\gamma = (\gamma_1, \dots, \gamma_p)$$

where $\gamma_j = 1$ if X_j is selected, and 0 otherwise.

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$$\gamma = (\gamma_1, \dots, \gamma_p)$$

where $\gamma_j = 1$ if X_j is selected, and 0 otherwise.

- Assign priors $f(\gamma)$, and also $f(\beta, \psi | \gamma)$. Interested in two things:
 - Posterior inclusion probabilities $\mathbb{P}[\gamma_j = 1 | \mathbf{y}]$ for variable X_j .
 - Posterior model probabilities $\mathbb{P}[\gamma = \gamma_k | \mathbf{y}]$ for model γ_k .

1 George and McCulloch's (1993) Stochastic Search Variable Selection [SSVS]

$$y_i = \alpha + \beta_1 X_{i,1} + \cdots + \beta_p X_{i,p} + \epsilon_i$$
$$\epsilon_i \sim N(0, \psi^{-1}) \text{ iid}$$

Priors on β and γ

$$\beta_j | \gamma_j \sim \gamma_j N(0, c_j^2 t_j^2) + (1 - \gamma_j) N(0, t_j^2)$$
$$\gamma_j \sim \text{Bern}(p_j)$$

- t_j and c_j are tuning parameters.
 - ▶ Suggested values are $(\text{SE}(\hat{\beta}_j)/t_j, c_j) = (1, 5), (1, 10), (10, 100),$ or $(10, 500)$.
 - ▶ $\text{SE}(\hat{\beta}_j) = \sqrt{\hat{\psi}^{-1}(\mathbf{X}^T \mathbf{X})_{jj}}$ under the full model.

2 Kuo and Mallick's (1998) sampler [KM]

$$y_i = \alpha + \gamma_1 \beta_1 X_{i,1} + \cdots + \gamma_p \beta_p X_{i,p} + \epsilon_i$$

$$\epsilon_i \sim N(0, \psi^{-1}) \text{ iid}$$

Priors on β and γ

$$\beta_j \sim N(b_j, d_j^2)$$

$$\gamma_j \sim \text{Bern}(p_j)$$

- Choices for b_j and d_j reflect prior beliefs on β .
- In the absence of prior information
 - ▶ Choose $b_j = 0$
 - ▶ Standardise the \mathbf{X} variables, and choose $d_j = d$ such that $1/2 \leq d \leq 4$

3 Dellaportas et. al. (2002) Gibbs Variable Selection [GVS]

$$y_i = \alpha + \gamma_1 \beta_1 X_{i,1} + \cdots + \gamma_p \beta_p X_{i,p} + \epsilon_i$$
$$\epsilon_i \sim N(0, \psi^{-1}) \text{ iid}$$

Priors on β and γ

$$\beta_j | \gamma_j \sim \gamma_j N(b_j, d_j^2) + (1 - \gamma_j) N(u_j, s_j^2)$$
$$\gamma_j \sim \text{Bern}(p_j)$$

- u_j and s_j are tuning parameters. Choices include
 - ▶ $u_j = \hat{\beta}_j$, the OLS estimates, and correspondingly $s_j^2 = \widehat{\text{Var}}(\hat{\beta}_j)$.
 - ▶ $u_j = 0$ and $s_j^2 \propto d_j^2$, but kept low.
- As before, we can choose $b_j = 0$ and $d_j = d$ with large d (after standardising \mathbf{X}) if no prior information.

Priors

- Priors for β_1, \dots, β_p

SSVS $\beta_j | \gamma_j \sim \gamma_j \mathcal{N}(0, 500^2 \cdot \widehat{\text{Var}}(\hat{\beta}_j)/10^2) + (1 - \gamma_j) \mathcal{N}(0, 500^2)$

KM $\beta_j \sim \mathcal{N}(0, 4^2)$

GVS $\beta_j | \gamma_j \sim \gamma_j \mathcal{N}(0, 10^2) + (1 - \gamma_j) \mathcal{N}(\hat{\beta}_j, \widehat{\text{Var}}(\hat{\beta}_j))$

- Priors for $\gamma_1, \dots, \gamma_p$

▶ $\gamma_j \sim \text{Bern}(1/2)$

▶ This shows our indifference between any choice of variables

- Priors for other parameters

▶ $\alpha \sim \mathcal{N}(0, 1000)$

▶ $\psi \sim \Gamma(0.001, 0.001)$

▶ Not too bothered about estimating these - just let the data take care of them

Simulated example

- Simple variable selection problem with $p = 5$ and $n = 50$.
 - ▶ Draw $\mathbf{X}_1, \dots, \mathbf{X}_5 \sim N(\mathbf{0}, \mathbf{I}_{50})$.
 - ▶ Generate response variables $\mathbf{Y} = \mathbf{X}_4 + \mathbf{X}_5 + \epsilon$.
 - ▶ ϵ drawn from $N(\mathbf{0}, 2^2 \mathbf{I}_{50})$.

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 - ▶ Generate response variables $\mathbf{Y} = \mathbf{X}_4 + \mathbf{X}_5 + \epsilon$.
 - ▶ ϵ drawn from $N(\mathbf{0}, 2^2 \mathbf{I}_{50})$.
- Simulation results for 10,000 MCMC samples

SSVS				KM			GVS		
	$\hat{P}[\gamma_j = 1 \mathbf{y}]$	S.E.		$\hat{P}[\gamma_j = 1 \mathbf{y}]$	S.E.		$\hat{P}[\gamma_j = 1 \mathbf{y}]$	S.E.	
γ_1	0.03	0.01		0.03	0.01		0.03	0.01	
γ_2	0.16	0.04		0.10	0.01		0.11	0.01	
γ_3	0.02	0.01		0.03	0.01		0.03	0.01	
γ_4	0.80	0.07		0.84	0.02		0.87	0.01	
γ_5	0.78	0.08		0.95	0.01		0.93	0.01	
Rank	Model	Prob.	Odds	Model	Prob.	Odds	Model	Prob.	Odds
1	$X_4 + X_5$	0.63	1.00	$X_4 + X_5$	0.72	1.00	$X_4 + X_5$	0.73	1.00
2	X_2	0.09	7.16	X_5	0.10	7.32	X_5	0.07	10.6
3	$X_2 + X_5$	0.04	18.4	X_4	0.08	7.78	$X_2 + X_4 + X_5$	0.04	18.0

① Introduction

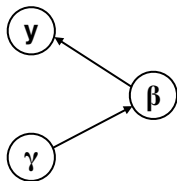
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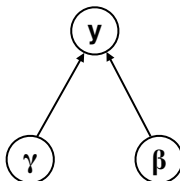
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⑤ Summary

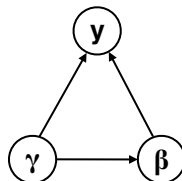
Comparison between the methods



$$f(\mathbf{y}|\beta)f(\beta|\gamma)f(\gamma)$$



$$f(\mathbf{y}|\gamma, \beta)f(\gamma)f(\beta)$$



$$f(\mathbf{y}|\gamma, \beta)f(\beta|\gamma)f(\gamma)$$

	SSVS	KM	GVS
Parameter space	Retains original	Does not retain original	
Tuning parameters	Many	None	Some
Priors for β	$\beta \gamma \sim N(\mathbf{0}, \mathbf{R}_\gamma \mathbf{D} \mathbf{R}_\gamma)$ $\mathbf{D} = \mathbf{I}_p$ $\mathbf{R}_\gamma = \text{diag}(a_j t_j)$ $a_j = (1 - \gamma_j) + \gamma_j c_j$	$\beta \sim N(\mathbf{0}, \mathbf{D})$ $\mathbf{D} = d^2 \mathbf{I}_p$	$\beta \gamma \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ $\boldsymbol{\mu}_j = (1 - \gamma_j) u_j$ $\boldsymbol{\Sigma}_{jk} = \gamma_j \gamma_k (d^2 \mathbf{I}_p)_{jk} + (1 - \gamma_j \gamma_k) \mathbf{1}_{[j=k]} s_j^2$

- All three are inefficient when there are strong correlations in the data.

Using I-priors in Bayesian variable selection

- Opportunity to use I-prior in each of the three methods. Simply replace $\mathbf{D} = \lambda^2 \psi \mathbf{X}^T \mathbf{X}$.
- The thought here is to replicate the correlations in the data into the prior covariance matrix of β .
- Which method works best with I-priors?
 - ▶ SSVS is unappealing due to the many tuning (hyper)parameters.
 - ▶ GVS is similar to KM, but designed to make the sampling more efficient. This was not seen in our simulations.
 - ▶ KM seems the simplest, “hands-free” method.

4 The KM I-prior model [I-prior]

$$\begin{aligned}y_i &= \alpha + \gamma_1 \beta_1 X_{i,1} + \cdots + \gamma_p \beta_p X_{i,p} + \epsilon_i \\ \epsilon_i &\sim \text{N}(0, \psi^{-1}) \\ i &= 1, \dots, n\end{aligned}$$

Priors

$$\begin{aligned}\beta &\sim \text{N}(\mathbf{0}, \lambda^2 \psi \mathbf{X}^T \mathbf{X}) \\ \gamma_1, \dots, \gamma_p &\sim \text{Bern}(1/2) \\ \alpha &\sim \text{N}(0, 1000) \\ \psi &\sim \Gamma(0.001, 0.001) \\ 1/\lambda^2 &\sim \Gamma(0.001, 0.001)\end{aligned}$$

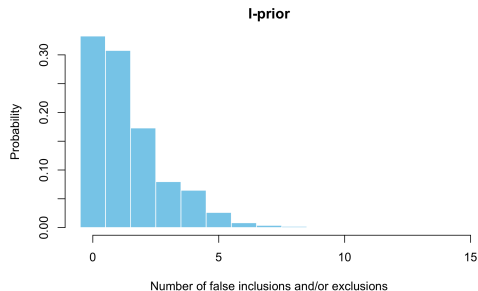
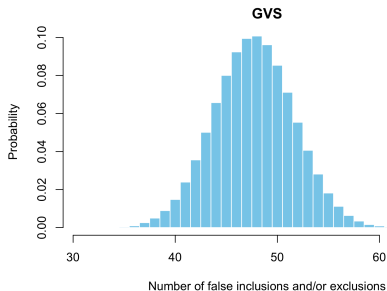
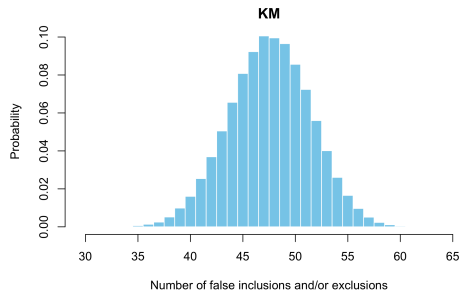
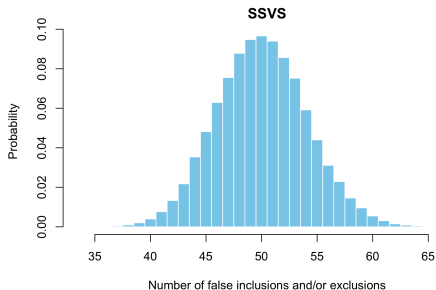
Simulation study

- Variable selection problem with $p = 100$ and $n = 150$ with artificial pairwise correlations between variables.
 - ▶ Draw $\mathbf{Z}_1, \dots, \mathbf{Z}_{100} \sim N(\mathbf{0}, \mathbf{I}_{150})$.
 - ▶ Draw $\mathbf{U} \sim N(\mathbf{0}, \mathbf{I}_{150})$.
 - ▶ Let $\mathbf{X}_j = \mathbf{Z}_j + \mathbf{U}$. This induces pairwise correlations of about 0.5.
 - ▶ Generate response variables $\mathbf{Y} = \mathbf{X}\beta_{true} + \epsilon$.
 - ▶ ϵ drawn from $N(\mathbf{0}, 2^2\mathbf{I}_{150})$.
- Let $\beta_{true} = (\beta_{-k}, \beta_k)$, where
 - ▶ $\beta_{-k} = (\beta_1, \dots, \beta_k) = (0, \dots, 0)$; and
 - ▶ $\beta_k = (\beta_{k+1}, \dots, \beta_{100}) = (1, \dots, 1)$.

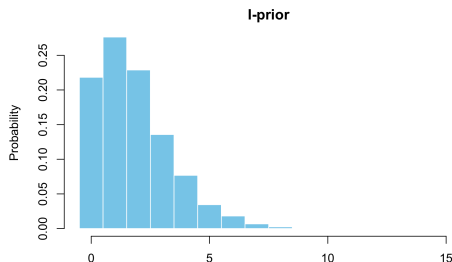
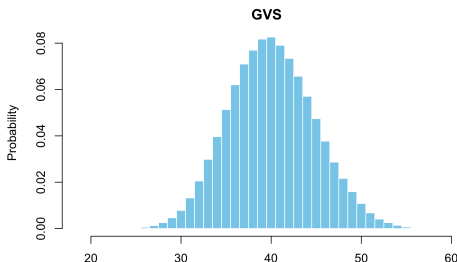
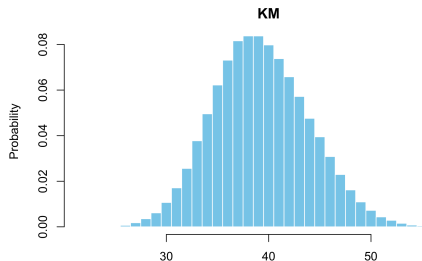
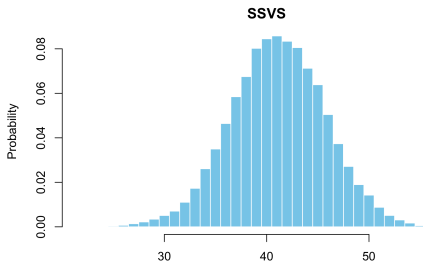
In other words, only variables X_{k+1} to X_{100} are used.

- The value of k is varied between 10, 25, 50, 75 and 90.
- 10,000 MCMC samples obtained for each scenario. Interested in how many false choices the models make. Each experiment repeated 10 times and results averaged.

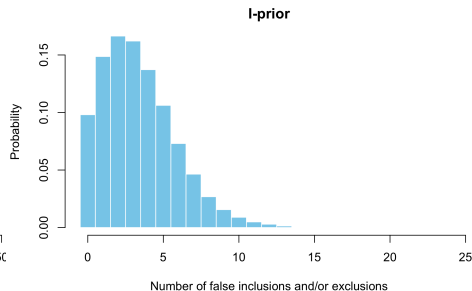
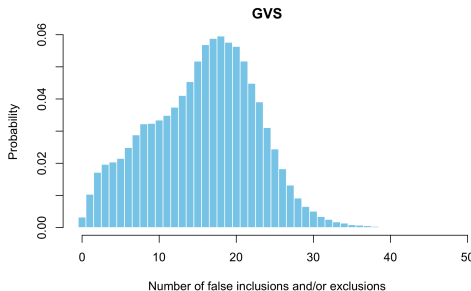
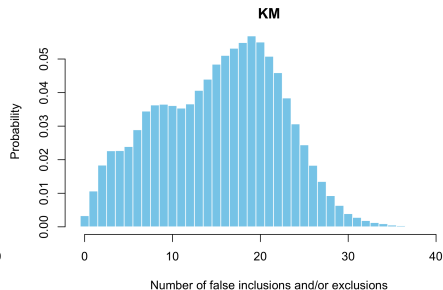
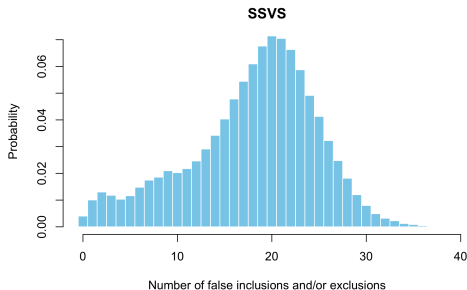
Scenario A (10,90)



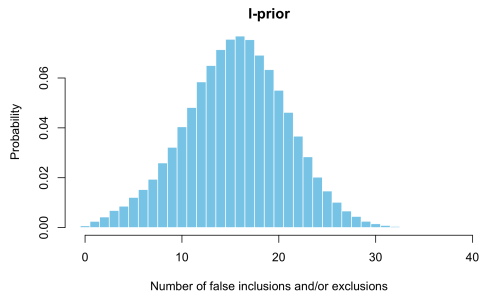
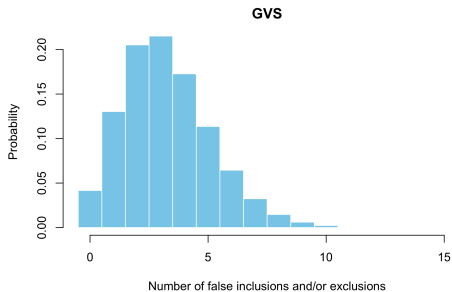
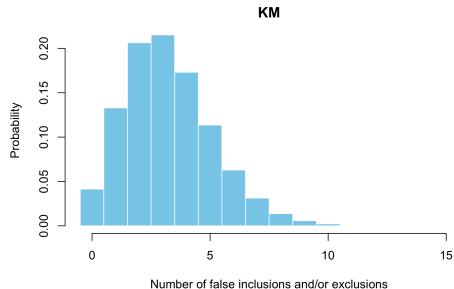
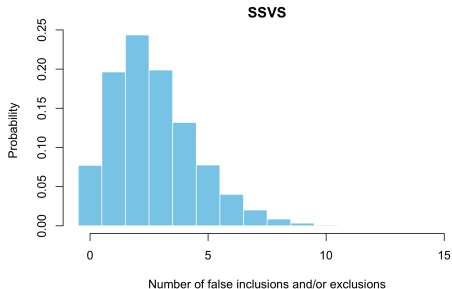
Scenario B (25,75)



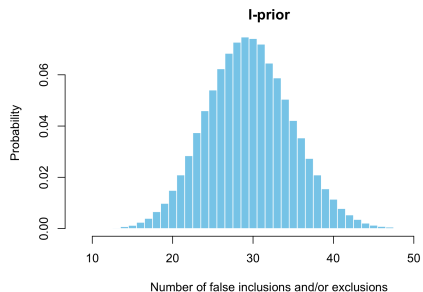
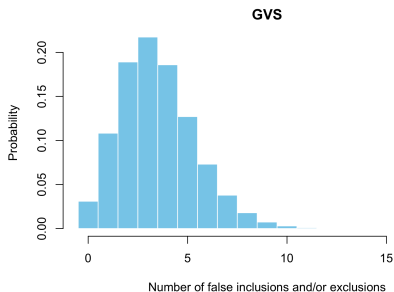
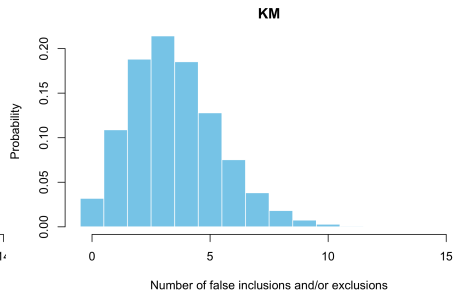
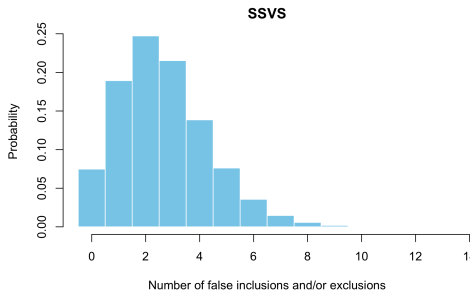
Scenario C (50,50)



Scenario D (75,25)



Scenario E (90,10)



Motivation for two-stage procedure

- I-priors performs very well when there are a lot of non-zero betas.
 - ▶ Strength comes from the Fisher information.
 - ▶ However, we can't expect I-priors to do well when few non-zero betas.
 - ▶ A lot of information becomes unnecessary and muddles the actual useful information.

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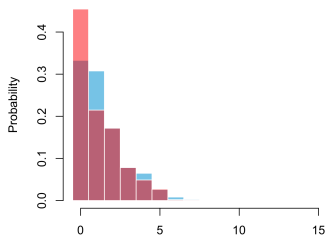
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- Two stage process
 - 1st Run the model. Keep only variables with posterior inclusion probabilities greater or equal to 0.5.
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- Barbieri and Berger (2004) showed that keeping such variables results in the most optimally predictive model being selected.

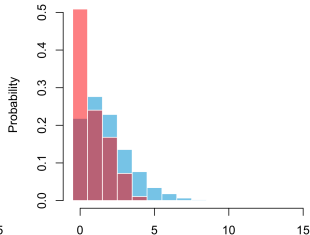
Two-stage simulation results

Scenario A (10,90)



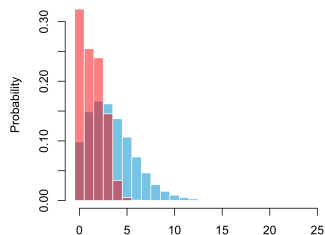
Number of false inclusions and/or exclusions

Scenario B (25,75)



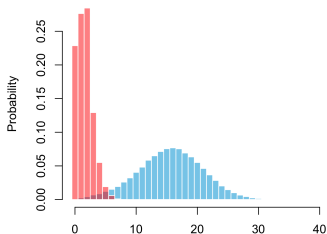
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Scenario C (50,50)



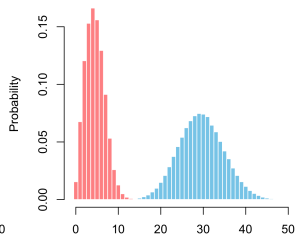
Number of false inclusions and/or exclusions

Scenario D (75,25)



Number of false inclusions and/or exclusions

Scenario E (90,10)



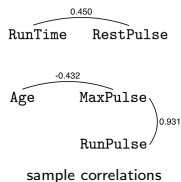
Number of false inclusions and/or exclusions

■ First stage
■ Second stage

Real world applications (1)

- Factors affecting aerobic fitness ($n = 30$ and $p = 6$) [Kuo and Mallick, 1998].
 - ▶ Response variable Oxygen, a measurement of oxygen uptake rate in mL/kg body weight per minute.
 - ▶ Covariates: Age, Weight, RunTime, RestPulse, RunPulse, MaxPulse.

	Full model	I-prior	Forward sel.	Back elim.
Intercept	104.2 (0.00)	80.8 (0.00)	103.3 (0.00)	98.6 (0.00)
Age	-0.24 (0.03)		-0.25 (0.02)	-0.21 (0.05)
Weight	-0.08 (0.15)		-0.08 (0.15)	
RunTime	-2.59 (0.00)	-2.97 (0.00)	-2.64 (0.00)	-2.75 (0.00)
RestPulse	-0.02 (0.72)			
RunPulse	-0.38 (0.00)	-0.38 (0.01)	-0.39 (0.00)	-0.36 (0.01)
MaxPulse	0.32 (0.03)	0.36 (0.02)	0.32 (0.03)	0.28 (0.05)
C_p	7.0	7.7	5.1	5.3
AIC	56.8	58.5	54.9	55.6
5-CV RMSE	2.59	2.71	2.50	2.54



Real world applications (2)

- Effects of air pollution on mortality in a US metropolitan area ($n = 60$ and $p = 15$) [McDonald and Schwing, 1973].
 - ▶ Response variable Mortality, a total age adjusted mortality rate.
 - ▶ Pollution potential data for HC, NO_x and SO₂.
 - ▶ Environmental considerations are Rain, JanTemp, JulTemp and Humid.
 - ▶ Socioeconomic considerations are Over65, Popn, Educ, Hous, Dens, NonW, WhiteCollar and Poor.

		Full model	I-prior	Min C_p	Back elim.
Environmental & demographic variables selected		All	Rain, JanTemp, JulTemp, Humid, Over65, Popn, Hous, NonW, Poor	Rain, JanTemp, JulTemp, Educ, NonW	JanTemp, Educ, NonW
Pollution effect	HC	✗	✗	✗	✓ $\beta = -0.98$
	NO _x	✗	✗	✗	✓ $\beta = 1.99$
	SO ₂	✗	✓ $\beta = 0.33$	✓ $\beta = 0.26$	✗
C_p		16.0	13.4	3.6	8.7
AIC		439.8	439.2	429.0	435.0
5-CV RMSE		50.6	41.6	37.1	38.6

Has this been done before...? A look at g-priors

- g-priors [Zellner, 1986] for linear regression coefficients has covariance matrix proportional to the inverse Fisher information

$$\beta \sim N(\mathbf{0}, g(\mathbf{X}^T \mathbf{X})^{-1})$$

- Popular choice of prior in Bayesian variable selection
 - ▶ “...use of $\propto (\mathbf{X}^T \mathbf{X})^{-1}$ tends to replicate design correlation” [George and McCulloch, 1993]
 - ▶ “The choice of $\propto (\mathbf{X}^T \mathbf{X})^{-1}$ serves to replicate the covariance structure of the likelihood” [Chipman et. al., 2001]
 - ▶ Used as an informative prior for variable selection problems, e.g. gene selection [Lee et. al., 2002]

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$$\epsilon \sim N(\mathbf{0}, \psi^{-1} \mathbf{I}_n)$$

$$\beta \sim N(\mathbf{0}, g(\mathbf{X}^T \mathbf{X})^{-1})$$

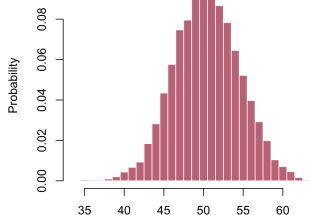
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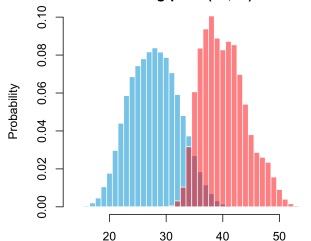
$$\left. \begin{aligned} \mathbf{y} &= \alpha + \mathbf{X}\beta + \epsilon \\ \epsilon &\sim N(\mathbf{0}, \psi^{-1}\mathbf{I}_n) \\ \beta &\sim N(\mathbf{0}, g(\mathbf{X}^T\mathbf{X})^{-1}) \end{aligned} \right\} \iff \begin{cases} \mathbf{y} = \alpha + \tilde{\mathbf{X}}\tilde{\beta} + \epsilon \\ \epsilon \sim N(\mathbf{0}, \psi^{-1}\mathbf{I}_n) \\ \tilde{\mathbf{X}} = \mathbf{X}(\mathbf{X}^T\mathbf{X})^{-1/2} \\ \tilde{\beta} = (\mathbf{X}^T\mathbf{X})^{1/2}\beta \\ \tilde{\beta} \sim N(\mathbf{0}, g^2\mathbf{I}) \end{cases}$$

g-prior results

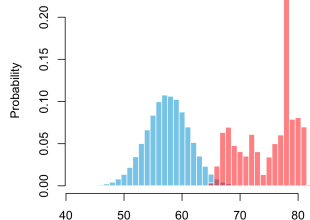
SSVS g-prior (10,90)



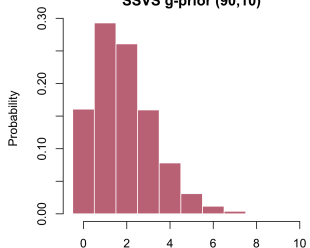
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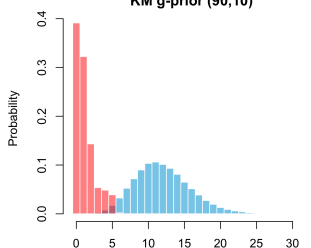
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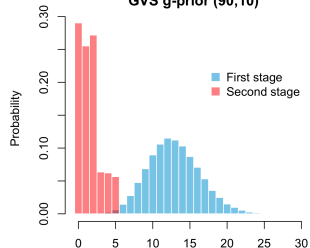
SSVS g-prior (90,10)



KM g-prior (90,10)



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① Introduction

② ASIDE: Regression modelling using I-priors

③ Bayesian variable selection

④ Using I-priors in Bayesian variable selection

⑤ Summary

Summary

- Variable selection under a Bayesian approach reduces to a problem of parameter estimation (i.e. γ).
- At the outset, wanted to find a simple and automatic way of running a variable selection model even in cases with (strong) collinearity.
- Our small contribution was to introduce an information theoretic prior in a two-stage approach. Good simulation results, but not very convincing in real world applications (yet).
- Things to do:
 - ▶ Find a way to accommodate individual scaling parameters for each variable - akin to original I-prior modelling.
 - ▶ Write own Gibbs sampling code for $p \ll n$ case.
 - ▶ Try this out on generalised linear models.

Some things I'd like to share

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 - ▶ 0%|=====|100%

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 - ▶ 0%|=====|100%
- If you work with matrices a lot, check out The Matrix Cookbook [Petersen and Pedersen, 2012].

End

Thank you!